Predicting Vehicular Travel Times by Modeling Heterogeneous Influences between Arterial Roads

Avinash Achar, Venkatesh Sarangan
Rohith Regikumar
TCS Research, IIT – Madras Research Park
Chennai 600113, INDIA

Anand Sivasubramaniam
Dept. of Comp. Sci & Eng.,
Pennsylvania State University,
State College, PA 16802, USA.

Abstract
Predicting travel times of vehicles in urban settings is a useful and tangible quantity of interest in the context of intelligent transportation systems. We address the problem of travel time prediction in arterial roads using data sampled from probe vehicles. There is only a limited literature on methods using data input from probe vehicles. The spatio-temporal dependencies captured by existing data driven approaches are either too detailed or very simplistic. We strike a balance of degrees of influence a given road may experience from its neighbors, while controlling the number of parameters to be learnt. Specifically, we use a NoisyOR conditional probability distribution (CPD) in conjunction with a dynamic Bayesian network (DBN) to model state transitions of various roads. We propose an efficient algorithm to learn model parameters. We also propose an algorithm for predicting travel times on trips of arbitrary durations. Using synthetic and real world data traces we demonstrate the superior performance of the proposed method under different traffic conditions.

1 Introduction

Travel-time prediction: Advances in affordable technologies for sensing and communication have allowed us to gather data about large distributed infrastructures such as road networks in real-time. The collected data is digested to generate information that is useful for commuters as well as the road network administrators. From the commuters’ perspective, travel time is perhaps the most useful information. Predicting travel time along various routes in advance with good accuracy allows commuters to plan their trips appropriately by identifying and avoiding congested roads. This can also aid traffic administrators to make crucial real-time decisions for mitigating prospective congestions, design infrastructure changes for better mobility and so on. Crowd-sourcing based applications such as Google Maps allow commuters to predict their travel times along multiple routes. While the prediction accuracy of such applications is reasonable in many instances, they may not be helpful for all vehicles. In certain countries, vehicles such as small commercial trucks are restricted to specific lanes with their own different (often lower) speed limit. Hence, the travel times and congestions seen by such vehicles could be different from the (possibly average) values that are predicted from crowd sourced applications. In such cases, customized travel-time prediction techniques are necessary.

Types of prediction models: Travel time prediction models can be broadly categorized into two types: traffic flow based and data-driven (Mori et al. 2015). The traffic flow models attempt to capture the physics of the traffic in varied levels of detail. They however suffer from important issues like need for calibration, being computationally expensive and rendering inaccurate predictions.

Data-driven models: The data driven models typically use statistical models which model traffic behavior to the extent required for the prediction at hand. They rely on real world data feeds for learning the parameters of the employed statistical model. A variety of data driven techniques to predict travel time have been proposed in the literature. Researchers have proposed techniques based on linear regression (Kwon, Coifman, and Bickel 2000; Nikovski et al. 2005), time-series models (Ishak and Al-Deek 2003; Vanajakshi, Subramanian, and Sivanandan 2009), neural networks (Li and Rose 2011), regression trees (Kwon, Coifman, and Bickel 2000; Nikovski et al. 2005) and bayesian networks (Hunter et al. 2009) to name a few.

Prediction in a freeway context (flow, travel time etc.) has been typically better studied compared to urban or arterial roads. This is because freeways are relatively well instrumented with sensors like loop detectors, AVI detectors and cameras. On the other hand, urban/arterial roads have been relatively less studied owing to complexities involved in handling traffic lights and intersections. Nevertheless, spread of GPS fixtures in vehicles/smart phones has rendered probe vehicle data a reasonable data source for arterial traffic (Liu, Yue, and Krishnan 2013; Aslam et al. 2012). Recently, DBN based approaches have been proposed to predict travel time on arterial roads based on sparse probe vehicle data (Herring et al. 2010; Hofleitner et al. 2012; Hofleitner, Herring, and Bayen 2012). Under real world traffic conditions, these various DBN techniques have been shown to significantly outperform other simpler methods such as time-series models.

Gaps and contributions: Current DBN based modeling approaches of congestion dependencies in road networks are either too meticulous to be used in large networks or too...
simplistic to be accurate. The modeling assumption in (Herring et al. 2010), albeit quite general, leads to an exponential number of model parameters. On the other hand, the model proposed in (Hofleitner et al. 2012) even though has a tractable number of parameters, assumes that the state of congestion in a given road is influenced equally by the state of congestion of all its neighbors, which can be pretty restrictive. In reality, different neighbors will exert different degrees of influence on a given road – for instance, the state of a downstream road which receives bulk of the traffic from an upstream road will exert a higher influence on the congestion state of the upstream road than other neighbors. In this paper, we propose a novel DBN based approach that models the individual influence of different neighbors while remaining computationally tractable. Our specific contributions include:

- We propose to use a ‘NoisyOR’ CPD for modeling the varying degrees of influence of different neighbors of a road. The degree of influence is offered as a separate parameter for each neighboring link. It also keeps the number of parameters to be learnt linear in the number of neighboring links.
- We develop a novel Expectation-Maximization (EM) based algorithm to learn the DBN parameters under the above NoisyOR CPD.
- We propose a new algorithm for predicting travel time of a generic trip that can span an arbitrary duration. Existing works can only handle trips that get completed within one DBN time step only.
- We test the usefulness of our approach on both synthetic data and real-world probe vehicle data obtained from the cities of (i) Porto, Portugal and (ii) San Fransisco. On synthetic data, relative absolute prediction error reduces by as much as 70% under the proposed method in the worst case. On real world data traces from Porto and San Fransisco, the proposed approach performs up to 14.6% and 16.8% better respectively than existing approaches in the worst case.

We note here that the proposed DBN with NoisyOR CPD transitions can be used in other domains as well, such as BioInformatics. Therefore, the proposed method has a wider reach than the specific transportation application discussed in detail in this paper.

2 Related work

Research based on probe vehicle data has been steadily on the rise of late given the wide spread use of GPS based sensing. Probe data has been utilized for various tasks like traffic volume and hot-spot estimation (Aslam et al. 2012), adaptive routing (Liu, Yue, and Krishnan 2013), and estimation and prediction of travel time. Travel time estimation is another (well studied) important task useful in particular for traffic managers. Since our focus in this paper is on prediction alone and since most of the travel estimation methods do not have predictive abilities, we do not elaborate on this further here. Please refer to App. B in (Achar et al. 2017) for a summary.

Literature on arterial travel time prediction using probe vehicles has been relatively sparse. We focus on DBN approaches which explicitly model the congestion state at each link. Among such DBN approaches, a hybrid approach that combines traffic flow theory and DBNs is proposed in (Hofleitner, Herring, and Bayen 2012). It captures flow conservation and uses traffic theory inspired travel time distributions. The state variables are no longer binary but the queue length built at each link. However, as discussed in (Hofleitner 2013) some of the model assumptions made in (Hofleitner, Herring, and Bayen 2012) like uniform arrivals are too strong and have limitations compared with physical reality. Reference (Hofleitner 2013) also recommends a relatively more data-driven approach as proposed in (Hofleitner et al. 2012) for real world adoption. Our proposed work is closely related to (Herring et al. 2010) and (Hofleitner et al. 2012) – we try to incorporate the best of these approaches while circumventing their drawbacks. In fact, the approach suggested in (Herring et al. 2010) may necessitate an exponential number of parameters to be learnt and hence suffers from overfitting.

3 DBN Model

Input Data: Probe vehicles are a sample of vehicles plying around the road network providing periodic information about their location, speed, path etc. Such vehicles act as a data source for observing the network’s condition. Such historical data is used for learning the underlying DBN model parameters. The learnt parameters along with current real-time probe data are used to perform short-term travel time predictions across the network. Real time is discretized into time bins (epochs or steps) of uniform size $\Delta$. At each time epoch $t$, we have a set of probe vehicle trajectory measurements. Each trajectory is specified by its start and end ($x_s$ and $x_e$) positions which come from successively sampled location co-ordinates, and sequence of links traversed in moving from $x_s$ to $x_e$. The data input to the algorithm is the set of all such trajectories collected over multiple time epochs. Notationally, for the $k^{th}$ vehicle’s trajectory at time step $t$, $x_{s,t}^k$ and $x_{e,t}^k$ are its start and end locations, and $L_t(k)$ is the sequence of links traversed. If $N_i^t$ denotes the number of active vehicles at time step $t$, then the index $k$ at time step $t$ can vary from 1 to $N_i^t$. Note that $N_i^t$ is a function of $t$ in general. In order to filter GPS noise and obtain path information, map matching and path-inference algorithms (Hunter et al. 2011) can be used. The notation used in this paper is summarized in App. A of (Achar et al. 2017).

DBN Structure

Fig. 1(a) shows the DBN structure (Herring et al. 2010; Hofleitner et al. 2012) that we use in this paper to capture spatio-temporal dependencies between links of the network. The arterial traffic is modeled as a discrete-time dynamical system. At each time step $t$, a link $i \in I$ in the network is assumed to be in one of two states namely, congested (1) or
uncongested (0). \( s^{i,t} \) denotes this state of congestion at link \( i \) and time \( t \). Note that these are hidden state variables. We denote by \( \pi_i \), the set of roads that are adjacent (both upstream and downstream) to road \( i \) including itself. The adjacency structure of the road network is utilized to obtain the transition structure of the DBN from time step \( t \) to time \( t + 1 \). Specifically, the state of a link \( i \) at time \( t + 1 \) is assumed to be a function of the state of all its neighbors \( \pi_i \) at time \( t \). In the DBN structure, this implies that the node corresponding to the link \( i \) at time \( t + 1 \) will have incoming edges from nodes in \( \pi_i \) at time \( t \).

We assume the travel time on a link to be a random variable whose distribution depends on the state of the respective link. The traversal time on a trajectory is a sum of random variables, each representing the traversal time of a (complete or partial) link of the trajectory. From the structure of the DBN (fig. 1(a)), given the state information of the underlying links, these link travel times \( (\tau^{i,t}) \), denoted as rectangles in fig. 1(a) are independent. Hence the conditional travel time on a path is a sum of independent random variables. In general, the first and last links in the set \( L_i(k) \) get partially traversed. In such cases, one can obtain the partial link travel times by scaling (linear or non-linear (Herring et al. 2010; Hofleitner et al. 2012)) the complete link travel time as per the distance. In this paper, we use linear scaling.

**Figure 1**: DBN structure and the proposed Transition CPD.

**Conditional probability distributions on DBN**

**Observation CPD**: Travel time distribution on a link \( i \) given its state \( s \), is assumed to be normally distributed with parameters \( \mu^{i,s} \) and \( \sigma^{i,s} \). We compactly refer to these observation parameters as \( (\mu, \sigma) \). The travel time measurement from the \( k^{th} \) vehicle at time epoch \( t \), \( y_{k,t} \) (denoted by circles in fig. 1(a)) is specified by the set of links traversed, \( L_i(k) \), and position of the start and end co-ordinates on the first and last links (namely \( x^{k}_{s,t} \) and \( x^{k}_{e,t} \) respectively). \( f(y_{k,t} | x^{k}_{s,t}, x^{k}_{e,t}) \) denotes the conditional distribution of a travel-time measurement, conditioned on the links traversed and start-end positions. Given state information of links along a path, owing to normality and conditional independence of these travel times, travel time on any path is also normally distributed. The associated mean and variances are sum of means and variances of individual link (possibly scaled) travel times.

**Existing Transition CPDs**: Let \( A(\eta^{i,t-1}, s^{i,t}) \) be the CPD that models influence exerted on road \( i \)'s state at time \( t \) by \( \eta^{i,t-1} \), the states of its neighbors at time \( t - 1 \). If this factor is a general tabular CPD as proposed in (Herring et al. 2010), then number of parameters grows exponentially with number of neighbors.

To circumvent this, (Hofleitner et al. 2012) chooses a CPD whose number of parameters is linear in the number of neighbors. Instead of a separate Bernoulli distribution for each realization of \( \eta^{i,t-1} \), it looks at the number of congested (or saturated) neighbors in the road network or parents in the DBN. Hence we refer to this as SatPat CPD in the rest of this paper. If \( a_{i,j} \) denotes the chance of congestion at the \( i^{th} \) link given exactly \( j \) of its neighbors are congested at the previous time instant, then

\[
A(\eta^{i,t-1}, s^{i,t}) = \prod_{j=0}^{b_{i,j}} \eta^{i,t-1} - s^{i,t} (1 - \eta^{i,t-1})^{j} = (1 - \eta^{i,t-1})^{b_{i,j} - j} \\
\]

where \( b_{i,j} \) is an indicator random variable which is 1 only when exactly \( j \) of link \( i \)'s neighbors are congested. As mentioned earlier, this CPD has a few shortcomings:

- It assumes all neighbors of a road have identical influence on a road's state. In particular it assumes an identical congestion probability (namely \( a_{i,j} \)) at \( i \) at time \( t \), given exactly one of its neighbors is congested at \( t - 1 \). This is irrespective of which of \( i \)'s neighbors is congested at \( t - 1 \).

- It is intuitive to expect that congestion probability of a road should increase with the number of congested neighbors. Specifically, one would expect that \( a_{i,0} \leq a_{i,1} \leq \cdots \leq a_{i,|\pi_i|} \). However, the learning strategy of (Hofleitner et al. 2012) doesn’t ensure this total ordering. Hence, it may be difficult to interpret real world dependency among neighboring roads from learnt parameters.

**Proposed Transition CPD**: To alleviate the above shortcomings, we propose to use a NoisyOR CPD (Koller and Friedman 2009) for modeling state transitions. If \( Y \in \{0, 1\} \), is the output and \( X = (X_1, X_2, \ldots X_n) \), \( X_k \in \{0, 1\} \), is the input, then the NoisyOR CPD is parameterized by \( n + 1 \) parameters, viz. \( (q_0, q_1, \ldots q_n) \), \( 0 \leq q_i \leq 1 \),
referred to as inhibitor probabilities. The CPD is given by:

\[ P(Y = 0 | X) = q_0 \prod_{k=0}^{n} q_k^{X_k}, \quad X_k \in \{0, 1\}. \]  

(2)

When \( q_0 = 1 \) and \( q_k = 0, \forall k > 0 \), we have the noiseless OR function. When one or more of the \( q_k \)s are non-zero, this CPD allows for a non-zero chance of the output becoming 0 in spite of one or more high inputs. In our context, with \( q_0 \) clamped to 1, each \( q_k \) exactly captures the chance of output being 0 (and hence the chance of congestion) when only the \( k^{th} \) neighbor is congested. Hence, the NoisyOR (unlike SatPat) captures influence of neighboring links in an independent and link-dependent fashion – with \( q_k \) representing the extent of influence from the \( k^{th} \) neighbor. As the number of congested inputs (neighbors) go up, the chance of unsaturation goes down as is evident from eq. 2. Hence it also captures the intuition of congestion probability increasing with the number of congested neighbors in the previous time step. The term \( (1 - q_0) \) captures the chance of congestion getting triggered spontaneously at a link (while all its neighbors are uncongested).

**Alternative representation for NoisyOR:** The NoisyOR comes from the class of ICI (Independent of Causal Influence) models (Heckerman and Breese 1994) and can be viewed as in Fig. 1(b). On each input line \( X_k \), there is a stochastic line failure function, whose output is \( Z_k \). The deterministic OR acts on the \( Z_k \)s. When the input \( X_k \) is zero, the line output \( Z_k \) is also zero. When \( X_k = 1 \), with inhibitor probability \( q_k \), line failure happens – in other words, \( Z_k \) is zero. The bias term \( q_0 \) controls the chance of the output being 1 in spite of all inputs being off. It is easy to check that CPD in Fig. 1(b) is given by eq. 2.

Under the NoisyOR CPD, the term which models the hidden state transitions can be expressed as \( A(\eta^{t-1}, \eta^{t-1}, s^{t-1}) \) where, \( \eta^{t-1} = \left[ \eta_i^{t-1}, \eta_j^{t-1}, \ldots, \eta_{|\mathcal{I}|}^{t-1} \right] \) with \( \eta_j^{t-1} \) representing the actual state of \( i^{th} \)'s neighbor \( j \) at time \( t-1 \). Similarly, \( \eta^{t-1} = \left[ \eta_0^{t-1}, \eta_1^{t-1}, \ldots, \eta_{|\mathcal{I}|}^{t-1} \right] \) with \( \eta_j^{t-1} \) denoting the new random variable introduced via the representation of Fig. 1(b). Note that \( \eta^{t-1} \) is of length \(|(\pi_i + 1)|\) while that of \( \eta^{t} \) is \(|\pi_i|\). Based on Fig. 1(b), we can write \( A(\eta^{t-1}, \eta^{t-1}, s^{t-1}) \), the transition factor, as follows:

\[
A(\eta^{t-1}, \eta^{t-1}, s^{t-1}) = P(s^{t-1}) P(\eta^{t-1} | \eta^{t-1}) \prod_{j=1}^{\pi_i} P(\eta_j^{t-1} | \eta_j^{t-1}) 
\]

(3)

\[
= \left[ (1 - q_{i,0}) P_{i,0}^{\eta_i^{t-1}} \prod_{j=1}^{\pi_i} \mathbb{1}_{\text{OR}(\eta_j^{t-1}) = s_i^{t-1}} \right] \prod_{j=1}^{\pi_i} \left[ (q_{i,j}) P_{i,0}^{\eta_j^{t-1}} (1-q_{i,j}) P_{i,j}^{\eta_j^{t-1}} \right] 
\]

where \( p_{i,j} = 1 - q_{i,j} \) and \( q_{i,j} \) is the probability that congestion at time step \( t - 1 \) in the \( j^{th} \) neighbor of link \( i \) does not influence \( i \) in time step \( t \). Similar to the SatPat CPD (eq.1), eq.3 demonstrates that a typical transition factor in the DBN under the NoisyOR CPD also belongs to the exponential family. This in turn makes M-step of EM learning feasible in closed form as explained later.

**Complete data likelihood under NoisyOR:** If \( s \) denotes the state of all links across all time, and \( y \) denotes the set of all travel time observations across all vehicles over time \( t = 1, \ldots, T \), the complete data likelihood is given by:

\[
p(s, y | \theta) = \prod_{t=2}^{T} A(\eta^{t-1}, \eta^{t-1}, s^{t-1}) \times \prod_{k=1}^{N_t} f(y_k | \mathcal{L}(k), t) \times \prod_{i \in \mathcal{I}} c^i(s^{t-1})
\]  

(4)

where \( c^i(0) \) is the marginal probability of link \( i \) being uncongested at time \( t \). We subsume this into \( A(\ldots) \) by constraining \( c^i(0) = q_i , 0 \). This is same as assuming all links start at time 0 uncongested. Note that, here \( \theta = (q, \mu, \sigma) \), where \( q \) refers to all NoisyOR parameters of each of the links.

### 4 Learning

An EM approach (refer App. C of (Achar et al. 2017) for details) is employed which is a standard iterative process involving two steps at each iteration. The E-step computes expectation of complete data log-likelihood (\( Q \)-function in short) at the current parameter values, while the M-step updates parameters by maximizing the \( Q \)-function. When the complete data log likelihood belongs to the exponential family, then learning gets simplified (Bishop 2006; Koller and Friedman 2009). The E-step involves just computing the Expected Sufficient Statistics (ESS). The M-step typically consists of evaluating an algebraic expression based on the closed form maximum likelihood estimate (MLE) under completely observable data, in which SS is replaced by ESS.

**E-step:** E-step which involves ESS computation, is actually performing inference on a belief network. Exact inference in multiply connected belief networks is known to be NP-hard (Cooper 1990). Since our DBN is also multiply connected with a large number of links, exact inference would lead to unreasonable run times. Hence we use a sampling based approximation algorithm for inference (Hofleitner et al. 2012). Specifically, we use a particle filtering approach. This involves storing and tracking a set of samples or particles. For each particle \( r \), we start off with a vector of uncongested initial states for all the links. At each time step \( t \), we grow each particle (sample) based on the current transition probability parameters (NoisyOR or SatPat). Each particle in state \( s^{t-1} \) is now weighted by \( \prod_{k=1}^{N_t} f(y_k | \mathcal{L}(k), t) \). An additional re-sampling of the particles based on these weights (normalized) is performed to avoid degeneracy. The required ESS (described above) are then estimated from these sample paths. As the name filtering indicates, the ESS at time \( t \) is calculated based on observations up to time \( t \), namely \( y^{t} \), rather than all observations. The ESS associated with observation parameters turns out to be \( P(s^{t-1} | \theta) = \mathbb{1}[y^{t} | \theta^{t}] \forall t, k, z \). Here, \( z \) refers to a binary vector of length \(|\mathcal{L}(k)|\).
M-step Update for DBN model

Observation updates: From eq. 4, it follows that Q-fn for the DBN model involves a sum of two terms: one exclusively a function of observation parameters ($\mu, \sigma$) and the other only of the transition parameters ($q$) for NoisyOR. Hence the joint maximization over ($\mu, \sigma$) and ($q$) gets decoupled. High time-resolution GPS observations are used to learn a 2-component Gaussian mixture at each link, which gives the means and variances of the individual link travel times. For convenient optimization, the variances thus obtained can be fixed and learning performed only over $\mu$ as carried out in (Hofleitner et al. 2012). However, one still needs iterative optimization owing to the complexity of the term involved.

Proposed transition parameters updates: Maximization of the second term involving hidden state transition parameters leads to an elegant closed-form estimate of the transition parameters for the proposed NoisyOR transitions. This is because each factor belongs to the exponential family.

Proposition 1. Given the observations $y$ and parameter estimate after the $\ell$th EM-iteration, $\theta^\ell$, the next set of transition parameters are obtained as follows:

$$q^{\ell+1}_{1,i} \propto \sum_{t=2}^{T} P(\eta^{i,t-1}_{j} = 1, \tilde{\eta}^{i,t-1}_{j} = 0 | y, \theta^\ell)$$  

$$p^{\ell+1}_{1,i} \propto \sum_{t=2}^{T} P(\eta^{i,t-1}_{j} = 1, \tilde{\eta}^{i,t-1}_{j} = 1 | y, \theta^\ell)$$  

where proportionality constants are same. Similarly for $j = 0$, the M-step updates are:

$$q^{\ell+1}_{0,i} \propto \sum_{t=1}^{T} P(\tilde{\eta}^{i,t-1} = 0 | y, \theta^\ell)$$  

$$p^{\ell+1}_{0,i} \propto \sum_{t=1}^{T} P(\tilde{\eta}^{i,t-1} = 1 | y, \theta^\ell)$$  

Please refer to App. D in (Achar et al. 2017) for a proof. The proof involves computing the $Q$-function for the proposed NoisyOR CPD and maximizing it in closed form. The above ESS are actually computed conditioned on $y^t$ (observations up to time $t$) and not $y$, via particle filtering as explained in the E-step. Attempting to do exact smoothing using all the observations would lead to unreasonable space complexities, given the large number of links. The above updates are for data observations from a single day. They can readily be extended to multiple days and handled efficiently in a parallel fashion as explained in App. E of (Achar et al. 2017). For a comparison of complexities between NoisyOR and SatPat, please refer to App. F in (Achar et al. 2017).

5 Prediction

Formally, given $\theta^*$ (learnt DBN parameters from historical data) and current probe vehicle observations up to time $t\Delta$ (or time bin $t$), the objective is to predict the travel time of a vehicle that traverses a specified trajectory (or path) $\Gamma = [i_1, i_2, \ldots i_{|\Gamma|}]$ starting at say $t\Delta$ (from time bin $(t+1)$).

Existing works (Herring et al. 2010; Hofleitner et al. 2012) estimate the travel time along $\Gamma$ under the assumption that it is lesser than $\Delta$ (or one time step). However, in general, the travel times for a trajectory can be greater than $\Delta$.

Challenge: As the DBN evolves every $\Delta$ time units, the state of the DBN estimated at $(t+1)^{th}$ time bin can be used to predict the network travel times only in the associated time interval $[t\Delta, (t+1)\Delta)$. If the given trajectory $\Gamma$ is not fully traversed by $(t+1)\Delta$, the DBN’s state has to be advanced to time epoch $t+2$. The estimated network state at $t+2$ should now be used to predict the network travel time in the interval $[(t+1)\Delta, (t+2)\Delta]$, and so on. In other words, the task of predicting the travel time along $\Gamma$ to contiguous trip segments $u_1, u_2, \ldots, u_M$, such that the expected travel time of segment $u_j$, $1 \leq j \leq (M-1)$, is $\Delta$ time units as predicted using the hidden state estimated at time epoch $(t+j)$; $u_M$ corresponds to the final trip segment in $\Gamma$ whose expected travel time is less than or equal to $\Delta$.

Approach: Without loss of generality, we assume that the end point $x_\Gamma$ of $\Gamma$ coincides with the end of link $i_{|\Gamma|}$. Algorithm 1 describes the procedure to predict the mean travel time, MTT, of $\Gamma$. App. G in (Achar et al. 2017) describes its correctness proof. CurStuff refers to the currently remaining suffix of $\Gamma$. CurSt is the fractional distance of the start point of CurStuff from the downstream intersection. $p^\ell_{M_i}$ - Expected travel time of traversing a $\ell$-length prefix segment, $L = [i'_1, i'_2, \ldots i'_\ell]$, of CurStuff. The idea is to first narrow down on the earliest $\ell$ (say $c$) at which $p^\ell_{M_i} > \Delta$. Subsequently, we need the exact position on $i'_\ell$ up to which expected travel time is exactly $\Delta$. The main component of the proof explained in App. G of (Achar et al. 2017) involves how to arrive at this exact position via a closed-form. This is utilized in lines 9 and 13 of Algorithm 1. FutStep keeps track of additional number of time steps until which particles are grown.

6 Experimental Results

Reference (Hofleitner et al. 2012), which proposes the SatPat CPD clearly demonstrates that a DBN with SatPat CPD outperforms baseline approaches based on time series ideas. Given this and comments made earlier in Sec. 2, we compare our proposed method with SatPat method only. We first test the efficacy of the methods on synthetic data. This is to better understand the maximum performance difference that can occur between the two approaches.

We implement learning by updating only the $q$ (NoisyOR case) or $a$ (SatPat case) parameters. During learning on synthetic data, we fix the observation parameters to the true values with which the data was generated. For ease of verification and since our contribution is in the M-update of hidden state transition parameters, we believe this suffices. However, it is straightforward to include $\mu$ as well in the iterative process as described in Sec. 4. The real data we consider in this paper is high time resolution probe vehicle data, where one can obtain independent samples of individual link times and learn a 2-component Gaussian mixture at each link. An-
other justification of our approach could be that there may not be a necessity to update $\mu$ and $\sigma$ once learnt via high time resolution GPS data.

We briefly summarize the synthetic data generation setup here and point the reader to App. I in (Achar et al. 2017) for additional details. The main idea is to use the DBN model of Sec. 3 with NoisyOR transitions and Gaussian travel times to generate trajectories. The generator takes as input a road network’s neighborhood structure and individual link lengths. The DBN structure is fixed from neighborhood information. The NoisyOR CPD gives a nice handle to embed a variety of congestion patterns. We choose CPD parameters to embed short-lived and long-lived congestions. The chosen synthetic network has 20 links with gridded one-way roads mimicking a typical downtown area. We chose 8 probe vehicles to circularly ply around the north-south region while another 8 along the east-west corridor.

Results on synthetic traces

We compare prediction error between proposed and existing methods as (true) trip duration is gradually increased. Specifically, we use the clearly distinct NoisyOR learning scheme (proposed) and SatPat learning scheme (existing) for comparison (Sec. 4). For prediction however, we emphasize that the algorithm used for comparisons here (for both NoisyOR and SatPat schemes) is not an existing algorithm but rather a generic one proposed here in Sec. 5 which can tackle trips of arbitrary duration. We randomly pick from the testing trajectories of each of the 16 probe vehicles, distinct non-overlapping trips of a fixed duration. We provide results of persisting (OR long-lived) congestion alone here. Results on short-lived congestion were found to be similar.

Each point in fig. 2(a) shows (Relative) Mean Absolute Error (MAE), obtained by averaging across all the distinct randomly chosen trips of a fixed duration (true trip time). By relative error, we mean the error divided by the true trip time. As true trip time (or prediction horizon) of the chosen trajectories is increased, the (MAE) also increases as intuitively expected. We also find that NoisyOR consistently gives more accurate predictions than SatPat justifying the need to model the varying influences of individual neighbors. For every prediction horizon, we also look for a trip on which difference in prediction errors between the proposed and existing approach is maximum. Fig. 2(b) gives the performance of both NoisyOR and SatPat with the maximum difference in prediction error, for a given prediction horizon. We see that prediction error difference can be as high as 70%, with NoisyOR being more accurate. Overall, it can be summarized that NoisyOR method’s predictions are significantly more accurate than the existing SatPat method. Further, NoisyOR learnt parameters can be interpreted better in real world than SatPat parameters.

---

**Algorithm 1:** Compute expected travel time of an arbitrary length query route

**Input:** $\theta^*$, Query Path $\Gamma = [i_1, i_2, \ldots, i_L]$, $\alpha_s$ - fractional distance of $x_s$ from downstream end of $i_1$.

**Output:** Mean travel time (MTT) of traversing $\Gamma = [i_1, i_2, \ldots, i_L]$, starting at $t\Delta$ from $x_s$ on $i_1$.

1. Initialize $\text{MTT} = 0$, $\text{CurSuff} = \Gamma$, $\text{CurSt} = \alpha_s$, $\text{FutStep} = 1$, $\mathcal{P}$ = Set of particles grown upto $t$;
2. while $\text{CurSuff} \neq \phi$ do
   3. Grow all particles in $\mathcal{P}$ by one step (either as per NoisyOR or SatPat transitions);
   4. $L := \ell$-length prefix path of $\text{CurSuff}$, say $[i'_1, i'_2, \ldots, i'_L]$. $b_{k-1} := \ell$-length binary representation of $(k - 1)$.
   5. $M_{\ell}(k) := \text{CurSt} * \mu_{i'_1, b_{k-1}(1) + \sum_{j=2}^{\ell} \mu_{i'_j, b_{k-1}(j)}}, p_{\ell}(k) := P(s_{t+\text{FutStep}} = b_{k-1}\mid y^t, \theta^*)$, $(2^\ell$-length vectors).
   6. if $\exists \text{ an } \ell \text{ s.t. } p_{\ell}^T M_{\ell} > \Delta$ then
      7. Compute the least $\ell$ (say c) using binary search (Use $\mathcal{P}$, the current set of particles to compute $p_{\ell}$);
      8. if $c > 1$ then
         9. $\text{CurSt} \leftarrow 1 - \{(\Delta - p_{\ell}^T M_{c^-})/p_{\ell}^T M_{c^+}\}$, where $M_{c^-}, M_{c^+}$ are $2^c$-length vectors;
         10. $M_{c^-}(k) := \text{CurSt} * \mu_{i'_1, b_{k-1}(1) + \sum_{j=2}^{c-1} \mu_{i'_j, b_{k-1}(j)}}$, $M_{c^+} = [\mu_{i^c2,0} \mu_{i^c1,1} \mu_{i^c0,0} \mu_{i^c1,1} \ldots \mu_{i^c0,0} \mu_{i^c1,1}]^T$.
      11. $\text{CurSuff} \leftarrow$ suffix of $\text{CurSuff}$ (from c); $\text{MTT} \leftarrow \text{MTT} + \Delta$; $\text{FutStep} \leftarrow \text{FutStep} + 1$
      12. else $\text{CurSt} \leftarrow \text{CurSt}(1 - (\Delta/p_{\ell}^T M_{c^+}))$; $\text{MTT} \leftarrow \text{MTT} + \Delta$; $\text{FutStep} \leftarrow \text{FutStep} + 1$
   13. else $\text{MTT} \leftarrow \text{MTT} + p_{\ell}^T M_{|\text{CurSuff}|}$; $\text{CurSuff} = \phi$;
17. return MTT
Results on real-world probe vehicle data

![Empirical CDF of Prediction Errors](a) Empirical CDF
![One Step Prediction Error](b) Error time-series

Figure 3: City of Porto: Test trajectory duration = $\Delta (5$ min).

![Empirical CDF of Prediction Errors](a) Empirical CDF
![Two Step Prediction Error](b) Sample of Prediction Errors

Figure 4: San Francisco: Test trajectory duration = $2\Delta (10$ min).

PORTO: To validate on real probe vehicle traces, we first used GPS logs of cars operating in the city of Porto, Portugal. The data was originally released for the ECML/PKDD data challenge 2015. Each trip entry consists of the start and end time, cab ID and a sequence of GPS co-ordinates sampled every 15 seconds. The GPS co-ordinates in the data are noisy as many of them map to a point outside the road network. The GPS noise was removed using heuristics such as mapping a noisy point to one or more nearest links on the road network. The observation parameters $\mu_{i,s}^{1,a}$ and $\sigma_{i,s}^{1,a}$ are learnt for each link $i$ using the high resolution (15 sec) measurements as performed in (Hofleitner et al. 2012). We fix observation parameters to these values and learn only the transition parameters.

We choose a connected region of the Porto map which was relatively abundant in car trajectories. This region consisted of roughly 100 links. App. J in (Achar et al. 2017) shows the actual region we narrowed down to. We chose a few second order neighbors (neighbor’s neighbor) too to better capture congestion propagation. It is likely that a congestion originating at an upstream neighbor of a short link might actually propagate to a downstream neighbor of the short link in question within $\Delta$ minutes. To account for this possibility, we add such second order neighbors (both upstream and downstream) to the list of original neighbors. We quantified short by links < 75m in length and pick $\Delta = 5$ min.

Trajectories from 4 p.m. to 9 p.m. were considered. One can expect the traffic conditions to be fairly stationary in this duration. The traffic patterns during a Friday evening can be very different from the other weekdays, which is why we treated Fridays separately. For sake of brevity, we discuss results obtained on Fridays alone. We trained on the best (in terms of the number of trajectories) 24 Fridays. Training was carried out using both the proposed NoisyOR and existing SatPat CPDs. We tested the learnt parameters on two Fridays.

Fig. 3 shows the performance of both the proposed and SatPat method on trajectories (with true trip time equal to $\Delta$ minutes) one time epoch ahead of the current set of observations. Given the sparse nature of the data obtained, we focused on testing trips of one $\Delta$ duration. Fig. 3(a) shows the empirical CDF of the absolute prediction errors (in %). The empirical CDF essentially gives an estimate of the range of errors both the methods experienced. A relative left shift of the NoisyOR CPD prediction errors indicate a relatively better performance compared to SatPat. We also observe from the errors that the NoisyOR method has a relative absolute error of about 5% lower than SatPat on an average and a relative absolute error of about 14.5% lower in the worst case. Figure 3(b) gives a sequence of (one-step) prediction errors for both methods across a few consecutive time ticks around which data was relatively dense to report meaningful predictions. Note that the worst case error of 14.5% was obtained at the 33rd time tick around which NoisyOR method continues to do better than SatPat.

SAN FRANCISCO: We also considered a similar taxi data from a region (please refer to App. J in (Achar et al. 2017) for a map view) of the bay area of San Francisco. Specifically, we considered trajectories of $2\Delta$ duration for testing from this data. We trained both the NoisyOR and SatPat models on about 11 days of data collected from this region of about 275 links in the evening. We present results in Fig. 4 for test trajectories of $2\Delta$ duration. As before, the empirical CDF given in Fig. 4(a) has a relative left-shift in the NoisyOR’s CDF, indicative of its better performance. Further, Fig. 4(b) gives the trajectory-wise prediction error comparison and an improvement of up to 16.8% was observed in the worst case and about 6% on an average.

This vindicates that the proposed technique of modeling influences of different roads in propagating traffic congestion can indeed be helpful. We also note that the worst case performance different between NoisyOR and SatPat is not as pronounced as in the synthetic traces. This could be attributed to the one of the following reasons: (i) the underlying congestion propagation characteristics may not be too much link dependent; (ii) even if the congestion propagation is link dependent, enough samples from probe vehicles may not be present in the available data logs.

7 Discussions and Conclusions

To conclude the paper, we proposed a balanced data driven approach to address the problem of travel time prediction in arterial roads using data from probe vehicles. We used a NoisyOR CPD in conjunction with a DBN to model the varying degrees of influence a given road may experience
from its neighbors. We also proposed an efficient algorithm to learn model parameters. We also proposed an algorithm for predicting travel times of trips of arbitrary duration. Using synthetic data traces, we quantify the accuracy of the proposed method to predict the travel times of arbitrary duration trips under various traffic conditions. With the proposed approach, the prediction error reduces by as much as 50 – 70% under certain conditions. We also tested the performance on traces of real data and found that the proposed approach fared better than the existing approaches. A possible future direction is to generalize the proposed approach to model road conditions using more than two states.

We believe that our NoisyOR based DBNs can be useful in other domains as well, as in Bio-Informatics. Inferring gene regulation networks (Karlebach and Shamir 2008) from gene expression data is a very important problem in Bio-informatics. Discovering the hidden excitatory/inhibitory interactions among the interacting genes is of interest here. DBN based approaches based on continuous hidden variables have been explored for this problem (Perrin et al. 2003). The NoisyOR based DBN and the associated learning algorithm introduced in this paper can be a viable alternative to infer the underlying gene interactions by employing a fully connected structure among the interacting genes. The learnt $q_{i,j}$ values can potentially indicate the strength of influence. We intend to explore this too in our future work.

**References**


